# **Discrete choice (1)**

**Applied Econometrics for Spatial Economics** 

## **Hans Koster**

Professor of Urban Economics and Real Estate







- <u>Introduction</u>
   The RUM framework
- 3. Value of time
- Multiple alternatives
   Summary

## **1. Introduction**





### 1. Introduction

- 2. The RUM framework
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### **1. Introduction**







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**1. Introduction** 

- If you have data at your disposal you may use these data to answer practical questions:
  - What factors influence the carrier's selection of a port?
  - Which mode do people prefer to travel from A to B?
  - Where do people want to live?



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- *Continuous* choice: estimate marginal changes in behaviour
  - E.g. "when fuel price increases by 10%, the demand for fuel will decrease by 2%"
  - Standard micro-economic theory applies

- Transport demand often has a discrete (binary) nature
  - Some *x* impacts a discrete *y*
  - Then use discrete choice methods



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- Discrete choice methods
  - <u>Dependent variable y<sub>i</sub> is discrete</u>
- Why not use OLS?
- Let's have the standard OLS equation  $y_i = \beta x_i + \epsilon_i$  (1) where *i* indexes the individual



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- OLS may be consistent for binary choice
  - But, y<sub>i</sub> (and therefore ε) is not normally distributed
- Horrace and Oaxaca (2006)
  - Leads to biased and inconsistent estimates if  $\hat{y}_i$  lies 'often' outside the [0,1] interval
  - I show in Clip #9 why that is an issue...

- OLS does not necessarily provides a link with economic theory
- Not suitable for multinomial choice

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## This week

- Learn about how to deal with discrete choices
- ... and stated choice experiments

• Plan:

Lecture #1:

- **Lecture #2:**
- Lecture #3:

Assignment:

The random utility framework
Estimating binary choice models
Estimating multinomial choice models
Stated vs. revealed preference data
Estimate the value of time



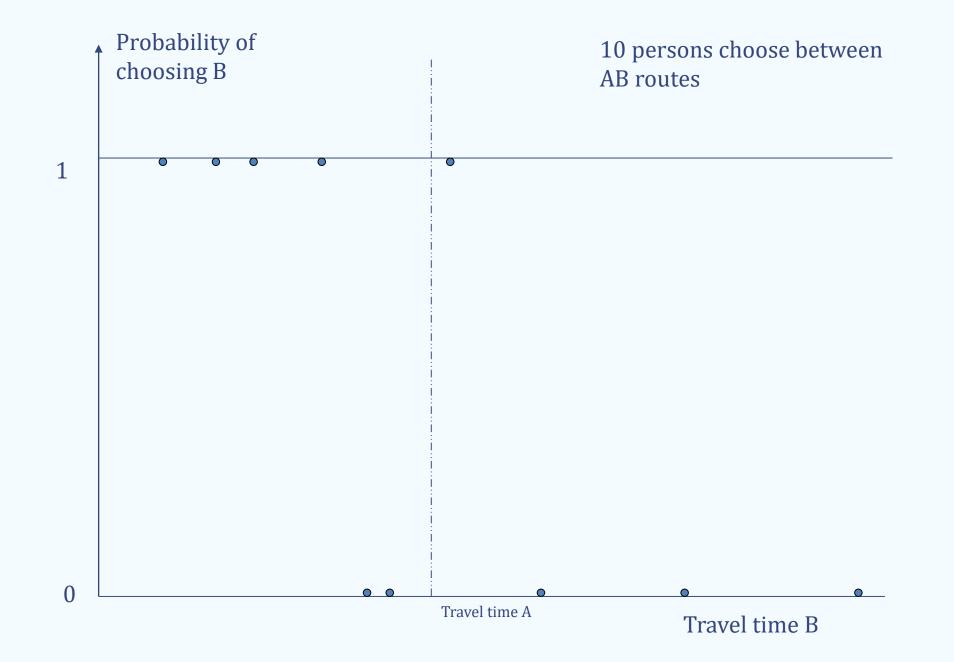
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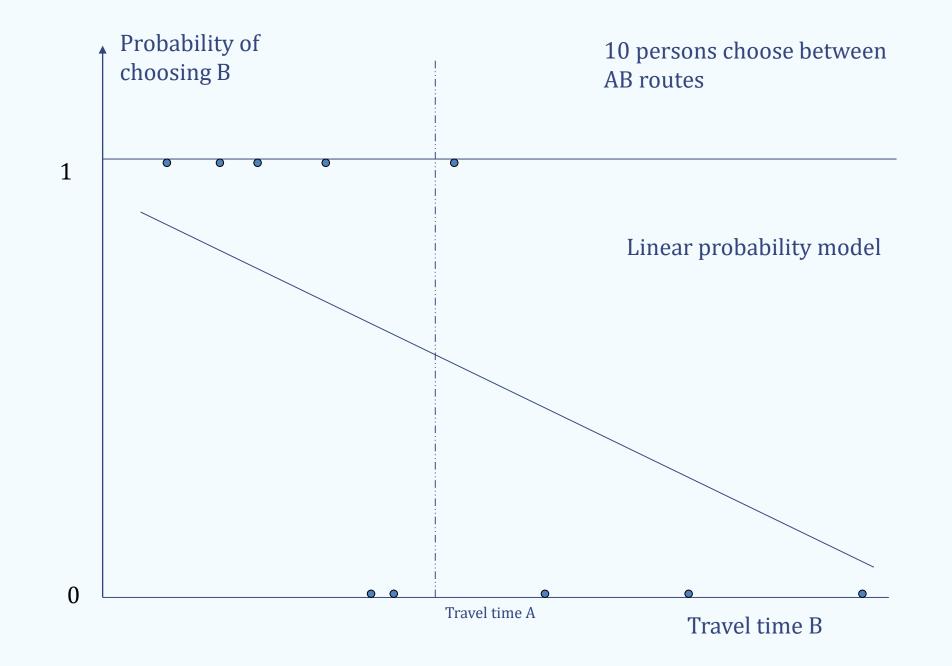
- Consider you have 10 individuals that choose between two routes A en B
- Travel time of A is 9 and of B is 10 minutes
- Some people take route *B* 
  - E.g. because they like particular features of , or they misjudge the travel time

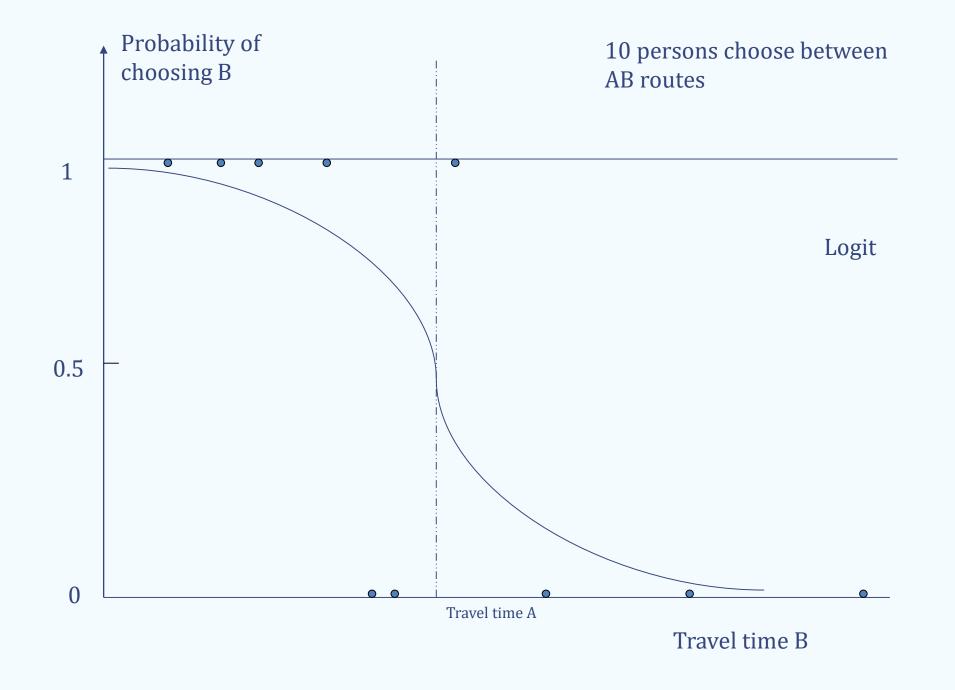
 Let's do a regression of whether or not you have chosen B on the difference between the travel time of A and B:

 $y_B = f(\alpha + \beta(travel time_B - travel time_A)) + \epsilon$ where  $y_B = 1$  if you choose *B* and zero otherwise









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- Indirect utility may be given by:  $U_{iA} = V_A(travel time_A) + \epsilon_{iA}$  (2)  $U_{iB} = V_B(travel time_B) + \epsilon_{iB}$  (3)
  - $V_A$ ,  $V_B \rightarrow$  deterministic utility

- Random terms:  $\epsilon_{iA}$ ,  $\epsilon_{iB}$ : <u>random taste variation</u>
  - Random utility model (RUM)
  - Note that the levels of  $U_{iA}$  and  $U_{iB}$  are not directly observed!

- $\Pr(Y = A) = \Pr(U_{iA} > U_{iB})$
- $\Pr(V_A + \epsilon_{iA} > V_B + \epsilon_{iB}) = \Pr(V_A V_B > \epsilon_{iB} \epsilon_{iA})$

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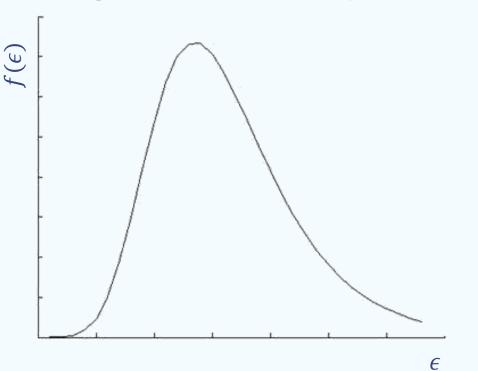
2. The random utility framework

- Two things are unknown
  - Which distribution for  $\epsilon$ 's?
  - What is the functional form for  $V_A$  and  $V_B$ ?



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- Which distribution for  $\epsilon$ 's?
  - $\epsilon$ 's are unobserved
  - You draw them from a distribution
  - Logit: Extreme Value Type I distribution





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2. The random utility framework

- Which distribution for  $\epsilon$ 's?
  - Extreme Value Type I distribution
  - Generates simple closed-form solutions!
    - $\rightarrow \Pr(V_A V_B > \epsilon_{iB} \epsilon_{iA})$
  - Daniel McFadden (1964)





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It appears that:

$$\Pr(Y = A) = \frac{e^{V_A}}{e^{V_A} + e^{V_B}}$$
(4)

• With two alternatives this can be written as:  $Pr(Y = A) = \frac{1}{1 + e^{V_B - V_A}}$ 



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- Which functional form for *V<sub>A</sub>* and *V<sub>B</sub>*?
  - Can be any function
  - Linear function is often assumed
  - Can be extended with multiple variables

$$U_{jA} = \beta p_{jA} + \kappa t_{jA} + \epsilon_{jA} \tag{5}$$

$$U_{jB} = \beta p_{jB} + \kappa t_{jB} + \epsilon_{jB} \tag{6}$$

where  $p_{jA}$  is the price of a trip and  $t_{jA}$  is travel time of alternative j

•  $\beta < 0, \kappa < 0$ 

Recall (from previous slide):

• 
$$\Pr(Y = A) = \frac{1}{1 + e^{\beta \left(p_{jB} - p_{jA}\right) + \kappa \left(t_{jB} - t_{jA}\right)}}$$



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- Important concept in Transport Economics: <u>Value of Time (VOT)</u>
  - "How much are you willing to pay to reduce your travel time with one hour, *holding utility constant*"

• Let's take the deterministic utility function  $U_{jA} = \beta p_{jA} + \kappa t_{jA} + \varepsilon_{jA}$ (7)



• When  $t_{jA}$  is measured in hours, the VOT can be written as  $\kappa/\beta$ 

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- <u>Value of time</u> is often used in cost benefit analyses
- VOT depends on trip purpose
  - Business €26.25/h
  - Commuting €9.25/h
  - Social purpose €7.50/h
- VOT depends on income
  - About 50% of net income



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• The choice probability for two alternatives:

 $\Pr(Y = A) = \frac{e^{\beta x_A}}{e^{\beta x_A} + e^{\beta x_B}}$ 

- Usually there are <u>more alternatives in the choice</u> <u>set</u>
  - Train, bus, car
  - Rotterdam, Antwerp, Hamburg
  - Routes to the VU

• Simply extend the logit formula:

$$\Pr(Y = A) = \frac{e^{\beta x_A}}{e^{\beta x_A} + e^{\beta x_B} + e^{\beta x_C}}$$



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- The <u>aggregate utility</u> derived from the choice set is summarised by the <u>logsum</u>:
  1
  - $E[CS] = \frac{1}{v} \ln(e^{\beta x_A} + e^{\beta x_B} + e^{\beta x_C})$
  - v is the marginal utility of income
  - Can be used in welfare estimates

- **Assume**  $\beta x_A = \beta x_B = 10$
- Now alternative *C* is added and  $\beta x_C = 1$
- The average utility per alternative decreases from 10 to 7 but E[CS] increase
  - 'Love of variety' effect

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4. Multiple alternatives

- Property of logit formula:
  - The *ratios* of choice probabilities for A and B do not depend on whether or not C is in the choice set
  - Independence of irrelevant alternatives

• 
$$\frac{\Pr(Y=A)}{\Pr(Y=B)} = \frac{\left(\frac{e^{\beta x_A}}{e^{\beta x_A} + e^{\beta x_B} + e^{\beta x_C}}\right)}{\left(\frac{e^{\beta x_B}}{e^{\beta x_A} + e^{\beta x_B} + e^{\beta x_C}}\right)} = \frac{e^{\beta x_A}}{e^{\beta x_B}}$$

• Let's find out whether this is a desirable property...



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- 'The Red Bus-Blue Bus' problem
- Choice set 1: Train, red bus, blue bus
- Assume market shares are 70, 15 and 15%

	<b>Train</b>	Red bus	Blue bus
V	2.54	1	1
Prob	0.700	0.150	0.150

## • Choice set 2: Train, red bus, so:

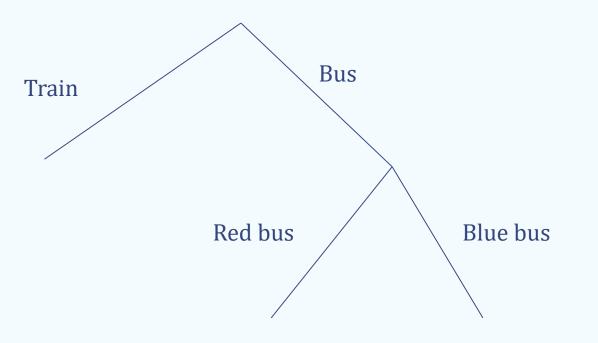
	Train	F	Red bus
V	2.5	54	1
Prob	0.82	23	0.177

- Probability to take the bus in choice set 2 is  $\frac{e^{1}}{e^{2.54}+e^{1}} = 0.177$ 
  - Higher probability not very realistic as red buses and blue buses are identical

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4. Multiple alternatives

- So, when some alternatives are more similar than other alternatives, the use of multinomial choice model may be misleading
- Use nested logit!





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4. Multiple alternatives

- Nested logit takes into account correlation between alternatives
  - But <u>define nests yourself</u>!



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Let us define utility as follows:  $U_{jg} = V_j + W_g + \epsilon_{jg}$   $V_j$  only differs within nests between alternatives j  $W_g$  only differs between nests g



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- We may write the probability to choose an alternative:
  - $\Pr(d_j = 1) = \Pr(g) \cdot \Pr(j \mid g)$

• 
$$\Pr(j \mid g) = \frac{e^{V_j / \lambda g}}{\sum_{k \in g} e^{V_k / \lambda g}}$$

• 
$$\Pr(g) = \frac{e^{W_g + \lambda_g I_g}}{\sum_{\tilde{g}} e^{W_{\tilde{g}} + \lambda_{\tilde{g}} I_{\tilde{g}}}}$$
  
with  $I_g = \log(\sum_{j \in g} e^{V_j / \lambda_g})$ 

- $\lambda_g = 1 \Rightarrow$  no correlation (multinomial logit)
- $\lambda_g \rightarrow 0 \Rightarrow$  perfect correlation (red bus/blue bus)
- Hence, when *j* and *k* are in the same nest:  $\frac{\Pr(d_j = 1)}{\Pr(d_k = 1)} = \frac{e^{W_g + V_j} / \lambda_g}{e^{W_g + V_k} / \lambda_g} = \frac{e^{W_g + V_j}}{e^{W_g + V_k}} = \frac{e^{V_j}}{e^{V_k}}$



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- So, nested logit probability depends on
  - Probability to choose a nest
  - Probability to choose an alternative within the nest

Note that Nested Logit does not imply a sequential choice



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**Today:** 

• How to deal with a binary dependent variable?

- Links to economic theory with random taste variation
  - Random utility model
  - Assume distribution of  $\epsilon_i$
  - Extreme Value Type I is convenient

Stated choice experiments can measure value of time



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## **5. Summary**

### **Tomorrow:**

- How to estimate binary choice models?
  - Use LPM, Logit or Probit

 Application to measure value of time, value schedule delay early and schedule delay late



# **Discrete choice (1)**

**Applied Econometrics for Spatial Economics** 

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# **Discrete choice (2)**

**Applied Econometrics for Spatial Economics** 

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Professor of Urban Economics and Real Estate







- 1. Introduction
- 2. Linear probability model
- 3. Logit
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### This week

- Learn about how to deal with discrete choices
- ... and stated choice experiments  $\bullet$

### **Plan:**

Lecture #2:

**Estimating binary choice models Estimating multinomial choice models** Lecture #3: Stated vs. revealed preference data Estimate the value of time **Assignment:** 



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- How to estimate binary discrete choice models?
- Three main options
  - 1. Linear probability model
  - 2. Logit
  - 3. Probit



- **1. Introduction**
- 2. <u>Linear probability model</u>
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• Regress 0/1 variable on characteristics of that choice and use OLS:  $Pr(d_j = 1) = \beta' x_j$ 

## Dataset example:

Chosen	Price	Time
1	14	12
0	25	5
0	15	15
1	15	13
1	4	45
1	3	40
0	20	10



- **1. Introduction**
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## **Advantages:**

• **Consistent when**  $0 \le \hat{y}_j \le 1 \forall j$ 

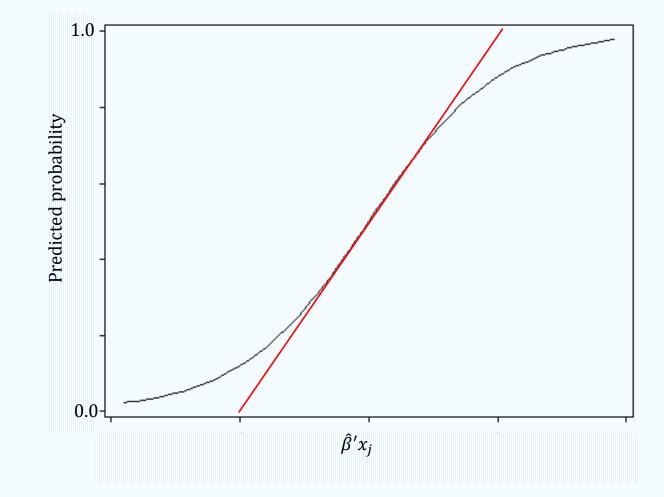


#### **1. Introduction**

- <u>Linear probability model</u>
   Logit
- 4. Probit
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### **Advantages:**

**Consistent when**  $0 \le \hat{y}_j \le 1 \forall j$ 





- **1. Introduction**
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# **Advantages:**

- **Consistent when**  $0 \le \hat{y}_j \le 1 \forall j$
- Easy to interpret
  - Say that β = -0.25 and x is measured in €, then for each euro increase in x, the probability to choose alternative *j* decreases by 25 percentage points

• 
$$\frac{\partial \Pr(d_j=1)}{\partial x} = \beta$$



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# **Advantages:**

- **Consistent when**  $0 \le \hat{y}_j \le 1 \forall j$
- Easy to interpret •  $\frac{\partial \Pr(d_j=1)}{\partial x} = \beta$

- Computationally feasible
  - Important for large panel datasets
- In practice, leads to very similar results as Logit and Probit



- **1. Introduction**
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## **Disadvantages:**

- No direct link with structural parameters of utility function
  - e.g. not able to calculate aggregate utility from choice set

- Biased for small samples and possibly inconsistent marginal effects
  - Linearity?

# Not suitable for multinomial choices



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Let's define

$$\Pr(d_j = 1) = \frac{1}{1 + e^{-\beta' x_j}}$$

Example: regress 0/1 variable on *differences* in characteristics of the alternatives

Chosen <sub>B</sub>	Price <sub>B</sub> -Price <sub>A</sub>	Time <sub>B</sub> -Time <sub>A</sub>
1	-14	5
0	5	0
0	15	-20
1	-8	13
1	-10	3
1	3	-5
0	20	10



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Recall

• 
$$\Pr(d_j = 1) = (1 + e^{-\beta' x_j})^{-1}$$

The change in the probability for one unit increase in x

3. Logit

$$\frac{\partial \Pr(d_j=1)}{\partial x_j} = \beta \frac{e^{-\beta' x_j}}{\left(1 + e^{-\beta' x_j}\right)^2}$$

- <u>Marginal effect depends on x<sub>j</sub>, so is not</u> <u>constant/linear</u>
  - For example, evaluate at mean values of *x*

Discrete choice (2)	3. Logit
<ol> <li>Introduction</li> <li>Linear probability model</li> <li>Logit</li> <li>Probit</li> <li>Application</li> <li>Summary</li> </ol>	• Marginal effects: • Use chain rule of differentiation • $\frac{\partial \Pr(d_j=1)}{\partial x_j} = -\left(1 + e^{-\beta' x_j}\right)^{-2} \times e^{-\beta' x_j} \times -\beta$

• 
$$\frac{\partial \Pr(d_j=1)}{\partial x_j} = \beta \frac{e^{-\beta' x_j}}{(1+e^{-\beta' x_j})^2}$$

- Dependent on x<sub>j</sub>, so is not constant/linear
  - For example, evaluate at mean values of *x*



- 1. Introduction
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# 3. Logit

# Software

- LOGIT **or** LOGISTIC **in STATA**
- REGRESSION BINARY LOGISTIC **in SPSS**
- In STATA you can select to report marginal effects
  - Use MARGINS after LOGIT command
  - Choose at which *x* the values are evaluated (e.g. at means)



Discrete choice (2)	3. Logit
<ol> <li>Introduction</li> <li>Linear probability model</li> <li>Logit</li> <li>Probit</li> <li>Application</li> <li>Summary</li> </ol>	Advantages of Logit: Predicted probability is always between one and zero
	<ul> <li>Clear link to random utility framework</li> <li>Log-sum may be used for welfare calculations</li> </ul>
	<ul> <li>Closed-form marginal effects</li> <li>Usually leads to very similar results as Probit</li> </ul>
	Can include 'fixed effects' (XTLOGIT in STATA)

• *e.g. to* control for individual heterogeneity



Discrete choice (2)	3. Logit
<ol> <li>Introduction</li> <li>Linear probability model</li> <li>Logit</li> <li>Probit</li> <li>Application</li> <li>Summary</li> </ol>	Disadvantages of Logit: ■ Why Extreme Value Type I distribution for <i>e</i> ?
	<ul> <li>Maximum likelihood / non linear model</li> </ul>



Discrete choice (2)		
<ol> <li>Introduction</li> <li>Linear probability model</li> <li>Logit</li> <li><u>Probit</u></li> <li>Application</li> <li>Summary</li> </ol>	•	W di
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- We may also assume that  $\epsilon_j$  is normally distributed, so  $\epsilon_j = N(0, \sigma^2)$ 
  - This implies  $Pr(d_j = 1) = \Phi(\beta' x_j)$
  - Central limit theorem?
  - However, no closed-form for cumulative normal distribution!

Marginal effects:

 $\frac{\partial \Pr(d_j=1)}{\partial x_j} = \beta \phi(\beta x_j)$ where  $\phi(\cdot)$  is the density function of the normal distribution

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Discrete choice (2)	4. Probit	
<ol> <li>Introduction</li> <li>Linear probability model</li> <li>Logit</li> <li><u>Probit</u></li> <li>Application</li> <li>Summary</li> </ol>	<ul> <li>Advantages:</li> <li>Normal distribution for ε<sub>j</sub> seems reasonable</li> <li> Central limit theorem</li> <li>Probability is always between one and zero</li> </ul>	

**Disadvantages:** 

- No closed-form marginal effects
- Hard to include many fixed effects



Discrete choice (2)	4. Probit
<ol> <li>Introduction</li> <li>Linear probability model</li> <li>Logit</li> <li><u>Probit</u></li> <li>Application</li> <li>Summary</li> </ol>	■ How to choose between the three models? • Probit estimates ≈ Logit estimates • Look at goodness of fit → Use $ d_j - \hat{d}_j $ • Check for robustness of marginal effects

- Large sample and interested in marginal effects?
  - → Usually linear probability model!
  - → There is an ongoing debate in economics on this issue



- **1. Introduction**
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#### 6. Summary

**Today:** 

- How to estimate binary choice models?
  - Use LPM, Logit or Probit

 Application to measure value of time, value schedule delay early and schedule delay late



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# **Tomorrow:**

Generalisations of logit models

**6.** Summary

- Multinomial logit
- Nested logit
- Conditional logit models
  - Poisson regression
- Data
  - Stated preference or revealed preference data



# **Discrete choice (2)**

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# **Discrete choice (3)**

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- 1. Introduction
- 2. Multinomial logit
- 3. Nested logit
- 4. Conditional logit
- 5. RP and SP data
- 6. Summary

# This week

- Learn about how to deal with discrete choices
- ... and stated choice experiments

# • Plan:

Lecture #1: The random utility framework
Lecture #2: Estimating binary choice models
Lecture #3: Estimating multinomial choice models
Stated vs. revealed preference data
Assignment: Estimate the value of time



- 1. Introduction
- Multinomial logit
   Nested logit
- 4. Conditional logit
   5. RP and SP data
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How to estimate these types of models? 

#### **Overview**

	# Alternatives	Coefficients
1. Binary Logit	2	Homogeneous
2. Multinomial Logit with alternative specific parameters	>2, <~10	Differ between alternatives
3. Nested Logit	>2, <~10	Usually homogeneous
4. Conditional Logit	>2	Homogeneous



- **1. Introduction**
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Recall:

$$Pr(Y = A) = \frac{e^{\beta x_A}}{e^{\beta x_A} + e^{\beta x_B} + e^{\beta x_C}}$$
  
But now let the coefficients be alternative-  
specific:

$$\Pr(Y = A) = \frac{e^{\beta_A x_A}}{e^{\beta_A x_A} + e^{\beta_B x_B} + e^{\beta_C x_C}}$$

- We cannot identify all the coefficients  $\beta_A$ ,  $\beta_B$ ,  $\beta_C$ , because we compare the results to a reference category
  - » Think of dummies
- Illustration: we can write the probability only in terms of differences with respect to one reference category, e.g.:

$$\Pr(Y = A) = \frac{1}{1 + e^{\beta_B x_B - \beta_A x_A} + e^{\beta_C x_C - \beta_A x_A}}$$

- **1. Introduction**
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- What demographic factors explain car ownership?
  - $0 = no \ car; \ 1 = 1 \ car; \ 2 \ge 1 \ car$
- **Data**, *n*=55958

respid	carown	hhsize	children	 socallow
100001	. 1	4	1	 0
100002	2	2	0	 0
100004	· 0	2	0	 1
100005	1	2	0	 0
100012	2	5	1	 0
622410	2	3	1	 0



- **1. Introduction**
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- 4. Conditional logit
- 5. RP and SP data
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- What demographic factors explain car ownership?
- Start with OLS
  - ... but car ownership is not really a continuous variable in the data

(Dependent variable: The number of cars in the household)		
	coeff.	s.e.
Household size	0.1745***	(0.0069)
Number of children in the household	-0.0045	(0.0163)
Social allowance (=1)	-0.6624***	(0.0120)
Male (=1)	0.1093***	(0.0051)
Age	-0.0031***	(0.0002)
Long term illness (=1)	-0.1317***	(0.0060)
Constant	0.7270***	(0.0152)
Number of observations	55,9	58
$R^2$	0.21	45

Table – EXPLAINING CAR OWNERSHIP

*Notes*: Robust standard errors are in parentheses. \*\*\* p < 0.01, \*\* p < 0.05, \* p < 0.10.

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#### MLOGIT YVAR XVARS, BASEOUTCOME(0)

Social allowance (=1)

Long term illness (=1)

Male (=1)

Constant

Age

 Table – EXPLAINING CAR OWNERSHIP

Outcome = 0	coeff.	s.e.
(base outcome)		
Outcome = 1	coeff.	s.e.
Household size	1.0039***	(0.0196)
Number of children in the household	-0.8290***	(0.0445)
Social allowance (=1)	-2.1039***	(0.0596)
Male (=1)	0.4870***	(0.0233)
Age	-0.0043***	(0.0007)
Long term illness (=1)	-0.3917***	(0.0248)
Constant	-0.6280***	(0.0506)
Outcome = 2	coeff.	s.e.
Household size	1.3039***	(0.0222)
Number of children in the household	-0.5191***	(0.0522)

-4.7724\*\*\*

0.5910\*\*\*

-0.0171\*\*\*

-0.7218\*\*\*

-1.9361\*\*\*

(0.2219)

(0.0295)

(0.0009)

(0.0358)

(0.647)

Number of observations	55,958
Log-likelihood	-48,268
Pseudo <i>R</i> <sup>2</sup>	0.1333

*Notes*: Standard errors are in parentheses. \*\*\* p < 0.01, \*\* p < 0.05, \* p < 0.10.

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- All the coefficients are compared to one base category!
- Coefficients are different for different alternatives
- Particularly useful when outcomes do not have a logical ordering
  - Bus, car, train
  - Holiday destinations
  - Otherwise: OLS or Ordered Logit

 If the number of alternatives is very large → too many coefficients to interpret meaningfully



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- Independence of irrelevant alternatives
  - Adding an alternative does not affect the relative odds between two other options considered
  - Solution: use Nested Logit
    - $\rightarrow\,$  Allows for correlation within nests

- Software
  - NLOGIT **in STATA**
  - Use Biogeme software
  - Limdep/nlogit



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- Often, <u>the number of alternatives is very large</u>
  - Location choice
  - Route choice
  - Holiday destinations
  - Choice of car
  - Partner choice
  - ....
- With Multinomial Logit this becomes infeasible
  - Unique coefficients for each alternative
  - Not necessary for large choice sets

<u>Conditional Logit</u>:

$$\Pr(d_j = 1) = \frac{e^{\beta' x_j}}{\sum_{k=1}^J e^{\beta' x_k}}$$

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- How to deal with large choice sets?
  - Number of observations in your regressions is number of alternatives × respondents

- 1. Model aggregate choices
- 2. Random selection of alternatives
- 3. Estimate count data models (Poisson)



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- 1. <u>Model aggregate choices</u>
- Modelling location choice
  - Focus on aggregate areas (municipalities)
- Choice of cars
  - Only distinguish between brands

 However, lack of detail makes results less credible



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- 2. <u>Random selection of alternatives</u>
- McFadden (1978)
  - Choose a random subset of *J* alternatives for each choice set, including the chosen option
  - This should not affect the *consistency* of the estimated parameters
  - Small-sample properties are yet unclear
- How large should / be?

- Applied in many good papers
  - e.g. Bayer et al. (2007, JPE)

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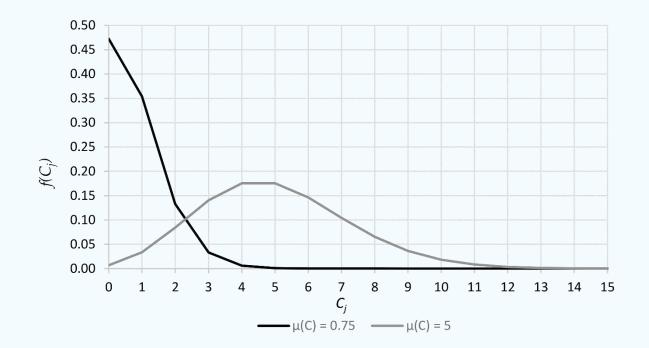
- 3. Estimate count data models
- Estimate Conditional Logit by means of a Poisson model

- A <u>Poisson regression</u> is a count data model
  - Dependent variable is integer
  - ... and should be Poisson distributed



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- 3. Estimate count data models
- Example of a Poisson distribution





• Equidispersion:  $\overline{y} = \sigma_y$ 

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- 3. Estimate count data models
- Estimate Conditional Logit by means of a Poisson model

- A <u>Poisson regression</u> is a count data model
  - Dependent variable is integer
  - ... and should be Poisson distributed
  - $C_j = e^{\beta' x_j} + \epsilon$

where  $C_j$  is the # of decision makers that have chosen a certain alternative

- Convenient interpretation of  $\beta$ 
  - When  $x_j$  increases with one,  $C_j$  increases with  $\beta \times 100$  percent

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- 3. Estimate count data models
- A Poisson model should give identical parameters to the Conditional Logit
  - Maximum likelihood functions are identical *up to a constant*
  - Guimarães et al. (2003)

- Hence, group observations based on their chosen alternatives
  - ... the number of firms choosing a certain location
  - ... the number of people buying a certain car



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- 3. Estimate count data models
- Implications
  - You cannot include characteristics of the decision maker (because you sum up all choices)!
  - Homogeneous parameters across the population

- Extensions
  - Include fixed effects
  - Negative binomial regression
  - Zero-inflated models
  - See Guimarães et al. (2004) for details



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# **Types of data**

- <u>Revealed preference (RP) data</u>
  - Observed or reported actual behaviour

- <u>Stated preference (SP) data</u>
  - Respondents are confronted with hypothetical choice sets

Combinations of RP and SP



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## Advantages of RP data

- Based on actual behaviour!!
- Use existing (large) data sources
  - Cheaper
  - No expensive experiments
- Panels of the same individuals over a long time



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- Lack of variability
- Collinearity (e.g. price and travel times)
- Lack of knowledge on the choice set
- Not possible with new choice alternatives
- Actual behaviour may not be first choice
  - University numerus fixus
- Perception errors and imperfect information
  - Airline tickets



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• Example of stated preference question

Suppose you have to ship a product from A to B			
Option 1		Option 2	
Price:	€ 1,000	Price:	€750
Handling time:	3 days	Handling time:	1 week
% does not arrive: 1.0%		% does not arrive: 1.3%	
What alternative will you choose?			



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# Advantages of SP data

- New alternatives
- New attributes
- Large variability is possible
- Problems of collinearity can be solved
  - 'Orthogonal design'
- Choice set is clearly defined



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- Information bias
- Starting point bias
- Hypothetical bias
- Strategic bias
- Errors



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- Information bias
  - The respondent has incorrect information on the context
  - Make your experiment as realistic as possible

# <u>Starting point bias</u>

- Respondents are influenced by the set of available responses to the experiment
- Test your design and choose realistic attribute values



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- Hypothetical bias
  - Individuals tend to respond differently to hypothetical scenarios than they do to the same scenarios in the real world.
  - Cognitive incongruity with actual behaviour
  - Again: make your experiment as realistic as possible
  - But otherwise hard to mitigate...

- Strategic bias
  - Respondent wants a specific outcome
  - (S)he fills in answers that are in line with desired outcomes



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- <u>Unintentional</u> biases
  - Information, starting point, hypothetical bias
- <u>Intentional</u> biases
  - Strategic bias

- Errors
  - Boredom
  - Respondents do not carefully read instructions
  - Respondents do not understand the questions



If there is good data available, I would prefer RP *(personal opinion)* 

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Today:

- Generalisations of logit models
  - Multinomial logit
  - Nested logit
  - Conditional logit
- Conditional Logit models can be estimated by count data models
  - Cannot include characteristics of the decision maker
- Data
  - Stated preference or revealed preference data



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Next week:

**Identification of causal effects** 

